

Weighted Integration Suggests That Visual and Tactile Signals Provide Independent Estimates About Duration

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Humans might possess either a single (amodal) internal clock or multiple clocks for different sensory modalities. Sensitivity could be improved by the provision of multiple signals. Such improvements can be predicted quantitatively, assuming estimates are combined by summation, a process described as optimal when summation is weighted in accordance with the variance associated with each of the initially independent estimates. This possibility was assessed for visual and tactile information regarding temporal intervals. In Experiment 1, 12 musicians and 12 nonmusicians judged durations of 300 and 600 ms, compared to test values spanning these standards. Bimodal precision increased relative to unimodal conditions, but not to the extent predicted by optimally weighted summation. In Experiment 2, 6 musicians and 6 other participants each judged 6 standards, ranging from 100 ms to 600 ms, with conflicting cues providing a measure of the weight assigned to each sensory modality. A weighted integration model best fitted these data, with musicians more likely to select near-optimal weights than nonmusicians. Overall, data were consistent with the existence of separate visual and tactile clock components at either the counter/integrator or memory stages. Independent estimates are passed to a decisional process, but not always combined in a statistically optimal fashion.

Public Significance Statement

We are able to judge the duration of events as they unfold (e.g., the time for which somebody holds our gaze). Sometimes, this information is conveyed to several of our senses at once (e.g., both seeing and feeling the duration of a caress). Theorists argue about whether time intervals are calculated separately for each sense or rely on a common centralized timer. This study suggests that when people experience the duration of events via both vision and touch, they gain a multisensory benefit, performing better than when they receive just visual or just tactile stimulation. This kind of benefit can only accrue if time is first estimated independently within each sense, suggesting that separate timers exist.

Keywords: time perception, interval timing, multimodal integration, audition, touch

Humans and other animals express their ability to time intervals through a wide variety of behaviors. In the lab, this ability is often assessed by requiring experimental participants to make judgments about the duration of events. However, our knowledge regarding the neurocognitive bases of these judgments remains hazy. One debate concerns the centralized versus distributed nature of the

hypothetical internal clock or clocks (Ivry & Schlerf, 2008; Ivry & Spencer, 2004). On the one hand, it is possible that all time-dependent behaviors rely on a single multimodal timer, through which a wide variety of sensory information is routed. On the other, it may be that there are different clocks for different purposes; for example, for sensory versus motor timing (Keele, Pokorny, Corcos, & Ivry, 1985; Marinovic & Arnold, 2012), implicit versus explicit timing (Coull & Nobre, 2008), short versus long interval timing (Lewis & Miall, 2003; Rammsayer, 1999), and for timing in different sensory modalities (Eijkman & Vendrik, 1965; Merchant, Zarco, & Prado, 2008).

The single versus multiple clock debate has received some attention from psychophysical studies focused on a single modality. For example, in both vision and touch, adaptation at a single spatial location can selectively affect subsequent interval judgments at that location, and not others (Johnston, Arnold, & Nishida, 2006; Watanabe, Amemiya, Nishida, & Johnston, 2010)

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perhaps implying multiple spatially localized timing mechanisms. However, there appears to be little or no statistical benefit from providing multiple visual inputs (which might reasonably be averaged to derive a more precise estimate of duration; see Ayhan, Revina, Bruno, & Johnston, 2012; and Morgan, Giora, & Solomon, 2008). This suggests either a severe attentional bottleneck for visual interval timing, or a single centralized mechanism that is modulated by localized visual inputs.

Further evidence pertinent to this debate comes from the multimodal literature. For example, the repetition of a visual stimulus with a particular duration can generate repulsive after effects, which can be measured when an adapted visual stimulus is compared to an auditory reference (and vice versa following auditory adaptation; see Heron et al., 2012). This implies that intervals in one modality either specifically or disproportionately adapt time perception within that same modality, suggesting modality-specific clocks. However, training on interval timing tasks in one modality can improve performance when precision is tested in a second sensory modality (Bartolo & Merchant, 2009; Bratzke, Seifried, & Ulrich, 2012; Nagarajan, Blake, Wright, Byl, & Merzenich, 1998; but see Lapid, Ulrich, & Rammsayer, 2009), instead suggesting a shared timing resource.

One behavioral method that can provide insights regarding the existence of multiple clocks involves assessing performance in unimodal conditions, and then seeking evidence for a bimodal improvement in precision resulting from the combination of initially independent sensory estimates (Macmillan & Creelman, 2005; Treisman, 1998). A widely adopted model of this process posits that information is combined via an optimally weighted summation process, known as maximum likelihood estimation (MLE) integration (Ernst & Banks, 2002; van Beers, Sittig, & Gon, 1999). In this case, each initially independent estimate receives a weighting in inverse proportion to its precision (i.e., the uncertainty regarding its accuracy). This model is typically contrasted with a less sophisticated strategy, in which observers rely on estimates arising from the modality that generally provides more precise information regarding the pertinent question. For example, auditory signals might dominate timing judgments, as auditory processing is typically more precise in the time domain than is vision.

The two computational strategies outlined above reflect just two possibilities. It is possible to envisage many intermediate models. For example, observers might be limited to using a single estimate, but may select this idiosyncratically, selecting their own most precise modality for a given experimental condition. Alternatively, they might make use of two modalities without actually averaging their estimates (e.g., via the max rule of probability summation from classic signal detection theory, in which both sensory estimates are evaluated and a decision is triggered if either reaches a criterion value). We could also consider an averaging of signals without optimal weighting of sensory evidence, or an averaging of signals containing partially correlated (rather than fully independent) sources of noise. In all cases, support for one or another of these models over others could be instructive regarding whether more than one independent estimate of duration has been derived on each experimental trial. Hence testing task performance in both unimodal and bimodal interval timing tasks could provide rich insights regarding the possible existence of multiple internal clocks.

To date, several studies have assessed cue combination for duration judgments via changes in bimodal precision.¹ Results have been mixed. The issue was, to our knowledge, first considered by Eijkman and Vendrik (1965). They applied detection-theoretic models to data from a filled duration discrimination task with auditory, visual, and audiovisual conditions. Observers attempted to detect deviations from a standard duration of 1 s. Bimodal performance was essentially identical to performance in the unimodal conditions (which were of similar difficulty to one another). The authors therefore concluded that noise had been perfectly correlated for unimodal signals, consistent with a single multimodal duration processor. This conclusion was at odds with the one they reached regarding the detection of intensity increments, where noise appeared uncorrelated (yielding bimodal enhancement). It also sits uncomfortably with findings pertaining to the summation of signals encoded by a single mechanism, such as when two spots of light are encoded by a single sensory detector—here sensitivity doubles, as if noise accrues only after signal combination.

More recent work has tended not to prioritize implications for models of the internal clock, focusing instead on establishing (or questioning) the general applicability of the MLE model of cue combination. For example, Burr, Banks, and Morrone (2009) found partial (but suboptimal) integration for empty intervals demarcated by audiovisual stimuli. A follow-up developmental study from the same group showed little evidence of integration (Gori, Sandini, & Burr, 2012). Around the same time, Van Wassenhove, Buonomano, Shimojo, and Shams (2008) investigated integration as part of a series of experiments focusing on the biasing effects of looming and receding stimuli for durations signaled by filled visual, auditory, and bimodal signals. The pertinent data (shown in their Figures S3 and S5) suggest partial, but suboptimal, facilitation in the bimodal case.

Some studies have tested relevant conditions without attempting to apply formal models of bimodal integration. For example, Gamache and Grondin (2010) illustrate a trend suggestive of multisensory facilitation in a subset of their unimodal and audiovisual conditions, which included matched auditory and visual standard durations. By contrast, a recent study testing high-functioning autistics and matched controls on an auditory, visual, and audiovisual interval comparison task obtained bimodal thresholds that were very similar to those of the best unimodal condition (Lambrechts, Yarrow, & Gaigg, 2016). A classic paper comparing auditory and visual time perception alongside audiovisual conditions (Walker & Scott, 1981) similarly showed no evidence for bimodal precision exceeding the best unimodal condition, a tendency we have also noted in several more recent publications (e.g., Rattat & Picard, 2012).

Two studies have used reproduction tasks and explicitly tested how auditory and tactile/motor (Shi, Ganzenmüller, & Müller, 2013) or visual and tactile (Tomassini, Gori, Burr, Sandini, & Morrone, 2011) time estimates are combined. These reported a partial integration and a very limited integration respectively (al-

¹ We focus on predictions regarding precision as these are generally considered to provide more compelling evidence of cue combination, specifically when bimodal precision exceeds that of the best contributing unimodal signal.

though it is difficult to properly partition sources of noise when using a reproduction task, in order to derive appropriate model predictions). However, Hartcher-O'Brien, Di Luca, and Ernst (2014) have recently reported optimally weighted integration for filled audiovisual stimuli of around 500 ms duration using an interval comparison task and have suggested that this should be generally obtained for filled-interval stimuli.

Given the mixed results, and the pertinence of cue combination studies to the debate regarding the architecture of the human timing system, further research investigating bimodal integration for duration judgments seems warranted. One issue particularly affecting audiovisual research on this topic is that experimenters have often used either clearly suprathreshold stimuli (in which case temporal sensitivity tends not to be well matched between auditory and visual stimuli, and predicted facilitation is therefore limited, reducing experimental power; e.g., Lambrechts et al., 2016) or have degraded stimuli by introducing masking noise (in which case it seems likely that the equalization of sensitivity that is achieved will reflect problems detecting temporal cues (presumably stimulus onsets and offsets) rather than a change in the scalar noise that usually dominates interval judgments; e.g., Hartcher-O'Brien et al., 2014). For both of these reasons, we chose to investigate integration of visual and tactile signals, which seemed more likely to give rise to similar levels of precision, even for nondegraded suprathreshold stimuli (Jones, Poliakoff, & Wells, 2009). We also investigated integration in individuals with differing levels of timing expertise, in order to determine whether integration might be experience dependent, operationalising this factor by assessing samples of participants with or without extensive musical training.

Experiment 1

Method

Participants. Twenty-seven participants were tested, 12 musicians (each of whom had received Associated Board of the Royal Schools of Music [ABRSM] Grade 8 qualifications) and 15 nonmusicians (with no prior experience of musical training). Three nonmusicians were excluded from the final analysis as they did not perform significantly above chance in one or more conditions (see data analysis, below). The final sample therefore consisted of 12 musicians, nine male, with a mean age of 26 years (from multiple disciplines of musical training: three percussionists, two brass players, and seven string instrumentalists), and 12 nonmusicians, four male, with a mean age of 23 years. Each participant provided informed consent, with testing procedures approved by the Department of Psychology Research Ethics Committee at City, University of London.

Apparatus and stimuli. The experiment was controlled by a Visual C++ program running on a PC interfaced with a 16 bit A/D card (National Instruments X-series PCIe-6323) generating digitized signals at 44.1 MHz. Visual stimuli were presented with a red LED (~0.5 degrees visual angle in diameter, ~60 mcd point source). It was placed on a desk to the left of a monitor approximately 50 cm from the participant. Vibrotactile stimuli were 200 Hz sine waves. They were delivered using a piezoelectric ceramic disk covered with a rubber sheath approximately 1 cm in diameter, powered by a bespoke amplifier, and were virtually silent. The

piezoelectric disk was pinched between the forefinger and thumb, which rested on the desk close to (~10 cm) but not obscuring the LED.

Design. A $2 \times 2 \times 3$ mixed factorial design included repeated measures on the last two factors. Factor 1 was the level of participant expertise: Musicians or nonmusicians. Factor 2 was the standard duration: 300 ms or 600 ms. The order of standards was randomized (without replacement) in each block. The final factor was stimulus modality: unimodal tactile, unimodal visual, and bimodal. This factor was blocked, and block order was counter-balanced across each set of six participants. In all, participants completed three blocks (600 trials).

Procedure. Participants sat at a desk with the computer screen and keyboard directly in front of them. The piezoelectric disk was held using the left hand in all conditions, including the unimodal visual condition. The right hand was used to indicate judgments using left or right arrow keys (or delete to cancel a trial due to an attentional lapse). Participants received 10 practice trials to familiarize themselves with the procedure in each condition, before beginning a block of 200 trials. No feedback on performance was given.

On each trial, a standard duration of either 300 ms or 600 ms was presented, followed by a 1,000 ms interstimulus interval that preceded the onset of a test stimulus (see Figure 1). The test stimulus was drawn at random from an adaptive distribution that was initially uniform (75–125% of the standard duration in 5% increments) but which could potentially expand to range from 25 to 175% of the standard duration: It was updated after each accepted trial according to the generalized Pólya urn procedure (Rosenberger & Grill, 1997) to sample the psychometric function in an efficient manner.

Following each test presentation, participants reported which interval had seemed longer, the first or second. We used this interval comparison task as it should provide clean estimates of sensory noise for the purposes of deriving model-based predictions about integration.²

Data analysis. For each participant, the proportion of tests judged longer than the standard was determined at each test duration in each condition. Data were imported into Matlab (The MathWorks Inc., Natick, MA) and maximum-likelihood fitted to cumulative Gaussian psychometric functions (with a fixed 1% lapse rate assumed). Points of subjective equality (PSE) and 84% thresholds (σ) were then estimated from these psychometric functions. PSE estimates represent the test value at which tests were judged to be longer than standard with a 0.5 probability. However, our main interest here was the threshold, which was estimated as the difference between test durations yielding “longer” judgments with probabilities 0.5 and 0.84. Under a standard observer model with Gaussian decision noise, this represents the standard deviation of the noise contributing to that decision.

Data were additionally fitted using a one-parameter horizontal line. A straight line would best fit data if a participant was randomly guessing with some unknown response bias for each

² As we noted earlier, it is challenging to properly partition sources of variance from some timing tasks, such as reproduction, for the purposes of fitting models of multisensory integration. We revisit this issue in the General Discussion.

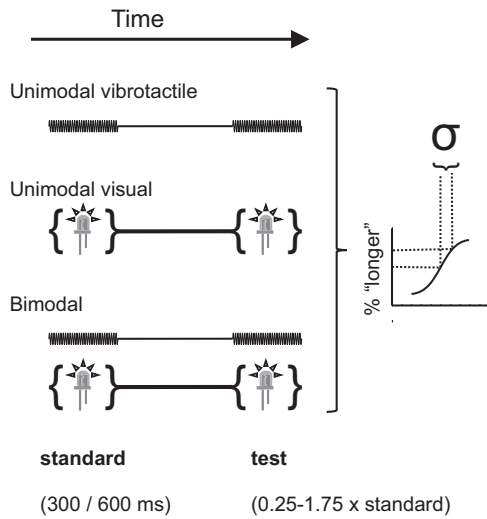


Figure 1. Schematic of methods for Experiment 1. Participants compared a standard interval to a test interval in unimodal tactile, unimodal visual, and bimodal conditions (in separate blocks). Standards could be 300 or 600 ms long (randomly interleaved) with tests spanning each standard. Fitted cumulative Gaussians provide threshold estimates corresponding to the standard deviation (σ) of the underlying difference distributions assumed to support duration judgments.

key. Participants were retained only when the cumulative Gaussian psychometric function fitted their data significantly better than a straight line ($p < .01$) in all conditions, assessed by comparing the difference in best-fitting model deviances against an appropriate chi-squared distribution (Wichmann & Hill, 2001).

Data from unimodal conditions were used to form predictions for both the optimally weighted integration (MLE) model and for a nonintegration model, assuming each observer relied entirely on the sensory signal producing their highest precision for that particular condition. For the MLE model, the prediction regarding the bimodal threshold (σ_{TV}) is well-established (Ernst & Banks, 2002):

$$\sigma_{TV}^2 = \frac{\sigma_T^2 \sigma_V^2}{(\sigma_T^2 + \sigma_V^2)} \quad (1)$$

Results and Discussion

Figure 2 shows mean threshold (σ) estimates for musicians and nonmusicians in all conditions, alongside model predictions. The inset graphs at the top of the figure present data from the unimodal tactile and visual conditions. Importantly, average performance was similar for vibrotactile and visual signals, providing scope for the combination of information to result in detectable performance improvements in bimodal conditions (relative to relying on one or other sensory modality alone). However, visual thresholds were generally slightly higher than tactile thresholds. As expected, unimodal thresholds increased near linearly with increasing standard duration (i.e., the scalar property of time) and were higher for nonmusicians than for musicians, consistent with many previous reports (e.g., Cicchini, Arrighi, Cecchetti, Giusti, & Burr, 2012; Rammsayer, Buttkus, & Altenmüller, 2012). These ob-

servations were confirmed by a $2 \times 2 \times 2$ mixed-design analysis of variance (ANOVA) showing a main effect of musical experience, $F(1, 22) = 14.93$, $p = .001$, a main effect of standard duration, $F(1, 22) = 51.85$, $p < .001$, and a main effect of sensory modality, $F(1, 22) = 4.55$, $p = .044$. A Group \times Standard Duration interaction, $F(1, 22) = 12.43$, $p = .002$, suggested that the increase in threshold with increasing standard duration was more pronounced for nonmusicians than for musicians. No other interactions reached significance.

Our main interest lay in determining whether performance in bimodal conditions reflected the combination of independent unimodal sources of information. The larger graphs toward the bottom of Figure 2 show predictions based on using the most precise unisensory signal (selected for each participant in each condition; black bars) versus predictions based on an optimally weighted combination of both unimodal signals (white bars). These predictions are plotted alongside mean thresholds for the bimodal conditions (gray bars). In general, bimodal thresholds lie between predictions based on optimal cue combination and those based on use of the best unimodal signal. There appears to be a trend for musicians to perform closer to the level predicted by optimal summation with a 600 ms standard compared to a 300 ms standard, with the opposite pattern emerging for nonmusicians.

To evaluate the degree to which each model could be rejected by the data, we ran a separate 2 (musical group) $\times 2$ (model vs. data) $\times 2$ (standard duration) mixed-design ANOVA for each model, with our interest focused on main effects and interactions involving the model versus data comparison. For the best unimodal model, there was a main effect of model versus data, $F(1, 22) = 10.18$, $p = .004$, suggesting that this model could be rejected, with no interactions involving this factor. Broadly the same pattern emerged for the optimally weighted integration model (model vs. data), $F(1, 22) = 6.46$, $p = .019$, allowing us to also reject that model. In this case, a three-way interaction, $F(1, 22) = 4.37$, $p = .048$, reflected the observed trend for musicians to differ from optimality more with a 300 ms standard ($t[11] = 2.68$, $p = .022$), whereas nonmusicians showed greater discrepancy for a 600 ms standard ($t[11] = 1.87$, $p = .088$).

Summarizing the results of Experiment 1, bimodal performance exceeded the predictions of a model assuming participants relied on their best unimodal estimate. However, bimodal precision failed to reach the level predicted by a model assuming an optimally weighted summation of unisensory estimates. Although our rejection of the best unisensory estimate model already provides some evidence against a neurocognitive architecture containing only a single amodal internal clock, we felt compelled to attempt to determine what decisional strategy or mechanism our observers were actually using. In particular, we were concerned that interval comparisons are likely to include subprocesses that each contribute noise to the decision, such that improvements in precision might reflect optimal integration for some, but not all sources of noise. This would clearly affect conclusions regarding any putative clock architecture. In particular, we wondered if participants might be integrating estimates of event onset and offset (likely affected by clock switch latency variability) without integrating estimates derived mainly from the integration of the intervening time (often thought of as a clock count), or vice versa.

To address this issue, we designed a second experiment in which six standard durations ranging from 100 to 600 ms were used. This

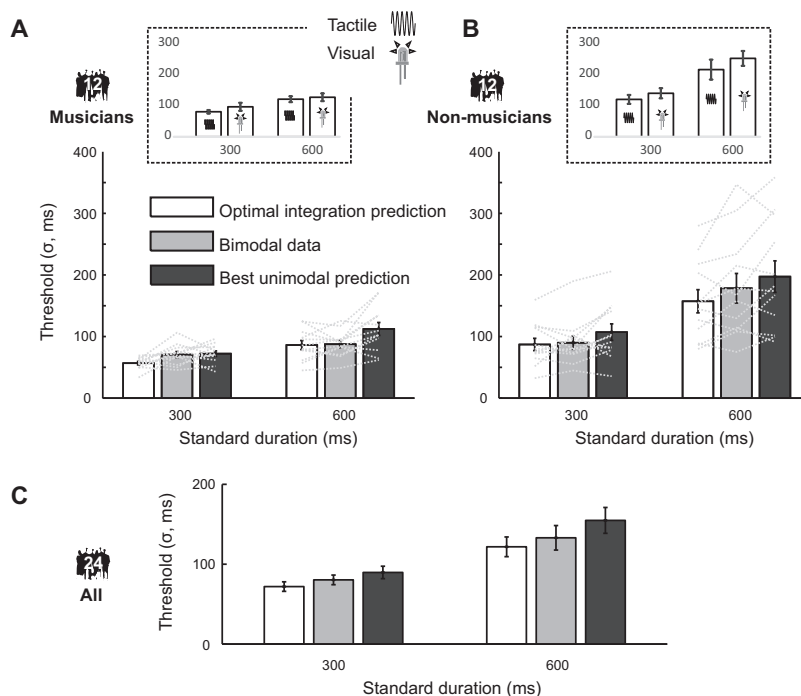


Figure 2. Results of Experiment 1 for (A) musicians, (B) nonmusicians, and (C) the entire group. Inset graphs (top) show mean threshold estimates for performance in the unimodal conditions. Main graphs (bottom) show threshold predictions for the bimodal condition, derived from the unimodal estimates by assuming either optimally weighted integration or reliance on the most precise sensory modality, alongside bimodal data. Individual participant thresholds are shown in light gray. Error bars denote standard error of the mean.

allowed us to apply a slope analysis method (Ivry & Hazeltine, 1995; Narkiewicz, Lambrechts, Eichelbaum, & Yarrow, 2015) to decompose judgment variability into a component contributed by nonscalar operations (including starting and stopping the clock) and a scalar component that grows with the duration being timed (reflecting counter and/or memory processes). With separate estimates, we were able to make predictions regarding the integration of one but not both sources of noise. We also took the opportunity to introduce a small discrepancy between tactile and visual components of our standard stimuli. This approach is commonly used (e.g., Ernst & Banks, 2002; Ley, Haggard, & Yarrow, 2009) to estimate the weight being assigned to each sensory modality and thus provide a further means of assessing the predictions of weighted integration models.

Experiment 2

Method

Methods were identical to those used in the first experiment, with the following exceptions.

Participants. Thirteen participants were tested including seven musicians (with ABRSM Grade 8 qualifications) and six psychophysical observers (i.e., lab members from local groups recruited specifically because they had experience participating in experimental tasks of this kind).³ One musician completed only a few blocks before being excluded from further participation, as

they reported great difficulty detecting the vibrotactile stimulus (confirmed by their very poor performance in tactile conditions). The final sample therefore consisted of six musicians (three pianists, two string instrumentalists, one brass player), all female, with a mean age of 22 years, and six psychophysical observers, four female, with a mean age of 29 years.

Design and procedure. Figure 3A schematises the design and procedure. Participants completed 18 blocks (3600 trials). A $2 \times 6 \times 3$ mixed factorial design was employed with repeated measures on the last two factors. Factor 1 reflected the characteristics of the participants: Musicians or psychophysical observers. Factor 2 varied the standard's duration: 100 to 600 ms. Standards were blocked with the following fixed order: 300, 400, 200, 500, 100, 600 ms. This ensured that conditions with the strongest implications for our slope analysis (see data analysis, below) were only completed after substantial practice. The third factor was modality

³ Of course, strictly, all of our participants could be considered psychophysical observers. We refer to this group as psychophysical observers, as opposed to nonmusicians, for two reasons. First, it contextualizes their strong timing performance; they were distinct from the musician group in another important respect, not just in being less musically expert. Second, we did not exclude participants from this group for having some degree of past musical training (although their musical attainment was in no sense equivalent to our musician group). In fact, in Experiment 2 we recruited musicians not so much because we expected them to be different, but rather because Experiment 1 had indicated that this group would perform timing tasks well.

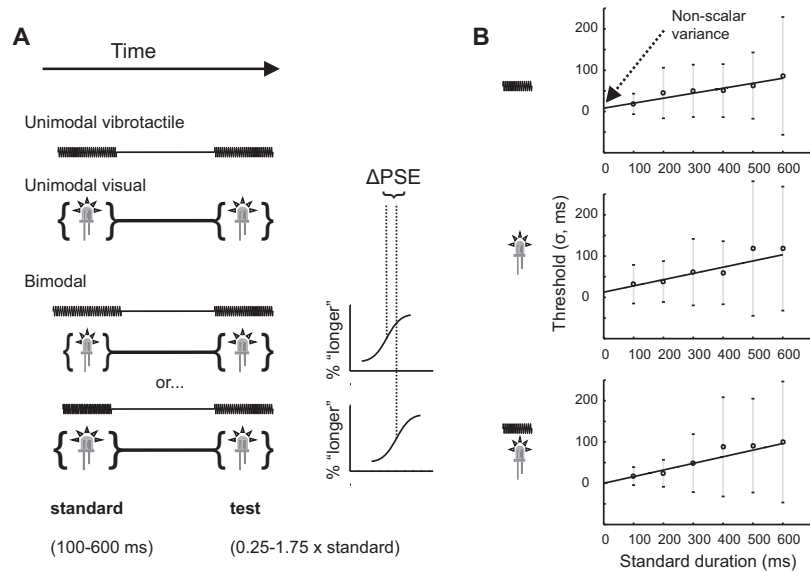


Figure 3. Schematic of methods for Experiment 2. **A:** Participants compared a standard interval to a test interval in unimodal tactile, unimodal visual, and bimodal conditions, with six standards ranging from 100 to 600 ms in duration (all in separate blocks). In bimodal conditions, there was a 5% discrepancy in duration between contributing unimodal stimuli (with the direction of this discrepancy randomly interleaved). Fits to tactile-long-visual-short versus visual-long-tactile-short trials were used to derive a shift in points of subjective equality (PSE) for the purposes of estimating the weights given to each modality. **B:** Scalar fits for one illustrative observer. Maximum-likelihood fits were obtained by using the standard error of the threshold parameter (obtained from separate cumulative Gaussian fits at each standard duration) to scale a Gaussian data model informing likelihoods. Error bars show the estimated standard error multiplied by 10 to better illustrate how precision of estimation scaled with standard duration.

of stimulus presentation: Unimodal tactile, unimodal visual, and bimodal. This factor was again blocked, with block order counterbalanced across each set of six participants, and a Latin square used to ensure each participant received a different order from all other participants in their group at each of the six standard durations.

The experiment was completed in sets of three blocks so that all three modalities were always presented within a single session for a given standard duration. A final nested manipulation was applied only to bimodal conditions, introducing a discrepancy between the tactile and visual components of the standard stimulus: tactile-long-visual-short versus visual-long-tactile-short. Discrepancies were introduced by presenting one modality for 97.5% of the standard duration, and the other for 102.5% of the standard duration, with midpoints aligned. This factor was randomized (without replacement) within each relevant experimental block, with 100 trials for each discrepancy per block of 200 trials.

Data analysis. Thresholds were determined for each participant at each of the 18 combinations of standard duration and sensory modality (collapsing across discrepancy conditions in the bimodal conditions). In addition to thresholds, we recorded 95% confidence intervals about thresholds to assist with model fitting (below).

A widely accepted property of interval timing is its adherence to a generalized version of Weber's law (Getty, 1975; Wearden & Lejeune, 2008). There is typically a linear relationship between interval duration and the standard deviation of trial-to-trial noise. This scalar noise rides atop a constant nonscalar variability, which can be found by determining the y-intercept of a straight-line

function fitted to threshold data (see Figure 3b). The six thresholds within each sensory modality were therefore fitted with a two-parameter model, with the intercept constrained to be ≥ 0 . To avoid undue influence from poorly estimated thresholds (and to appropriately weight the more precisely estimated thresholds typically observed at lower standard durations), we used a maximum-likelihood (rather than least squares) fit, assuming a Gaussian data model with separate scale parameters reflecting uncertainty for each threshold (based on their 95% confidence intervals). In addition to an estimate of nonscalar variability, this fitting procedure provided a set of cleaned data points (i.e., the scalar model's predicted values) which were used in all subsequent calculations and analyses.⁴

For this experiment, we estimated thresholds in all conditions, in addition to the weights assigned to the tactile modality in bimodal conditions. Empirical weights were derived by fitting cumulative Gaussians to data collapsed across all six standard durations (with

⁴ We consider these cleaned data preferable to the raw data as they make best use of well-established properties of time perception to reduce measurement noise. We return to this issue in the discussion. Note, however, that it would have also been difficult to have proceeded from the raw data in our analyses, as two of our 12 participants produced large and very poorly estimated thresholds in a small subset of conditions, despite generally performing well. With 18 conditions per participant, we preferred not to reject entire participants on the basis of one or two poor estimates, and our fitting procedure made this feasible.

the test values normalized by their standard durations and expressed as percentages) but separated on the basis of the discrepancy within the standard stimulus (tactile-long-visual-short vs. visual-long-tactile-short). The PSEs from these fits were differenced, and this difference scaled to generate an empirical estimate of tactile weight:

$$W_T = 0.5 + \frac{PSE_{TLVS} - PSE_{VLTS}}{10} \quad (2)$$

These empirical weights were used in modeling (see below) and also to test the optimally weighted integration model. For the MLE model, the prediction regarding optimal weights is well-established (Ernst & Banks, 2002):

$$W_T = \frac{\sigma_V^2}{\sigma_V^2 + \sigma_T^2} \quad (3)$$

For the purposes of this prediction, we used the average of thresholds at each standard duration, normalized by their respective standards.

Data from the unimodal conditions were also used to form predictions for six different models regarding bimodal thresholds. In addition to the optimally weighted integration and best unimodal predictions (outlined in Experiment 1), we calculated predictions for two hybrid models, in which nonscalar variance was integrated optimally and scalar variance came from the best unimodal stimulus, or vice versa. We also considered a model utilizing the max rule of probability summation: Observers were assumed to have access to both unimodal estimates, and to compare tests to standards in both cases, responding longer if either unimodal test exceeded its corresponding standard. Finally, we tested a model of weighted integration (i.e., averaging) but with weights based not on the optimal selection strategy described in Equation 3, but rather derived empirically (as described in Equation 2). Predictions for these final two models were based on Matlab simulations, with simulated data subjected to a cumulative Gaussian fit (in a similar manner to our real data) to generate predicted thresholds (cf. Treisman, 1998).

Our main inferential analysis again used ANOVA, but to assist data visualization and inference at the level of the individual participant, we also made use of nonparametric bootstrapping, using the bias-corrected method with 1999 bootstrap resamples. For bootstrap inference, we calculated 95% confidence intervals on the difference between estimates from pairs of conditions.

Results and Discussion

In Experiment 2 we included a small discrepancy between the vibrotactile and visual components of our bimodal standards. This allowed us to estimate the degree to which each participant relied on tactile (vs. visual) information to inform their decisions. Estimated tactile weights are plotted in Figure 4a, against weights that would be predicted for optimally weighted cue combination (based on performance in unimodal conditions). Across the sample as a whole, we did not obtain the predicted linear relationship (with a slope of 1.0) between optimal and empirical tactile weights ($r = .186, p > .05$). However, it is apparent that our musician subgroup (black circles) showed a much greater degree of optimality in their empirically derived tactile weights than did our psychophysical observers (white circles).

The additional standard durations and empirical estimates of tactile weights available in Experiment 2 allowed us to consider a wider range of model predictions regarding bimodal thresholds. Figure 4b plots the predictions of several models (in the leftmost column of graphs) for the group as a whole, and also separately for each subgroup. It is clear that for these participants, relying on the most precise (i.e., best) unisensory modality would generate very similar predictions to using this same strategy supplemented by the optimally weighted integration of only nonscalar sources of independent information (e.g., clock onset/offset times) or indeed to making use of both modalities independently via a max response rule (i.e., probability summation).⁵ Meanwhile, predictions from the optimally weighted integration model are very close to those of a model in which only scalar sources of independent information are combined via optimal weighting. Finally, weighted combination based on empirically derived tactile weights falls midway between these groups of predictions, at least for the sample as a whole. However, this form of cue combination provides near-optimal predictions about threshold for the musician subset (as they appeared to select near-optimal weights) but predicts thresholds similar to using the best unimodal stimulus for the nonmusician subset (who appeared to give an inappropriately high weighting to the vibrotactile stimulus).

Given the grouping of model predictions we observed, we went on to test whether just three of our candidate models differed from the data. As in Experiment 1, we did this by contrasting bimodal thresholds with the predictions of each model in turn, via ANOVA, focusing on effects of model versus data. In Figure 4, the set of graphs second from the left show model predictions alongside bimodal data for a model in which participants relied on the best unimodal stimulus in each condition of the experiment. A 2 (musical group) \times 2 (model vs. data) \times 6 (standard duration) mixed-model ANOVA (with Greenhouse-Geisser corrections for violations of sphericity) showed a significant main effect of model, $F(1, 10) = 5.33, p = .045$, but also a three-way interaction, $F(5, 50) = 9.62, p = .010$. The interaction reflected the fact that for the psychophysical observers, bimodal thresholds did not differ from best unimodal model predictions,⁶ whereas for musicians, they did, $F(1, 5) = 7.80, p = .039$, with a trend toward a greater difference at longer durations (Model vs. Data \times Duration Interaction), $F(5, 25) = 6.49, p = .050$.

The row of graphs second from the right show model predictions alongside bimodal data for predictions based on the optimally weighted combination of cues. Here, it is apparent that the musicians match the data but the psychophysical observers do not. ANOVA revealed a main effect of model, $F(1, 10) = 13.71, p = .004$, but also a three-way interaction, $F(5, 50) = 7.14, p = .023$. In this case, the interaction was driven by a difference between

⁵ In its simplest form this approach also predicts large shifts in PSE for bimodal conditions (Treisman, 1998). We have not reported PSEs here but did not observe shifts of this kind. In classic signal detection-theoretic experiments (which use a single stimulus level) probability summation outperforms single-modality detection, but this does not always hold true when a full psychometric function is mapped, in part because this function becomes highly asymmetric when modalities differ greatly in precision, so cumulative Gaussian fits are poor and generate large threshold estimates.

⁶ There was an interaction between the model versus data comparison and the standard duration, but no pairwise comparison approached even uncorrected significance.

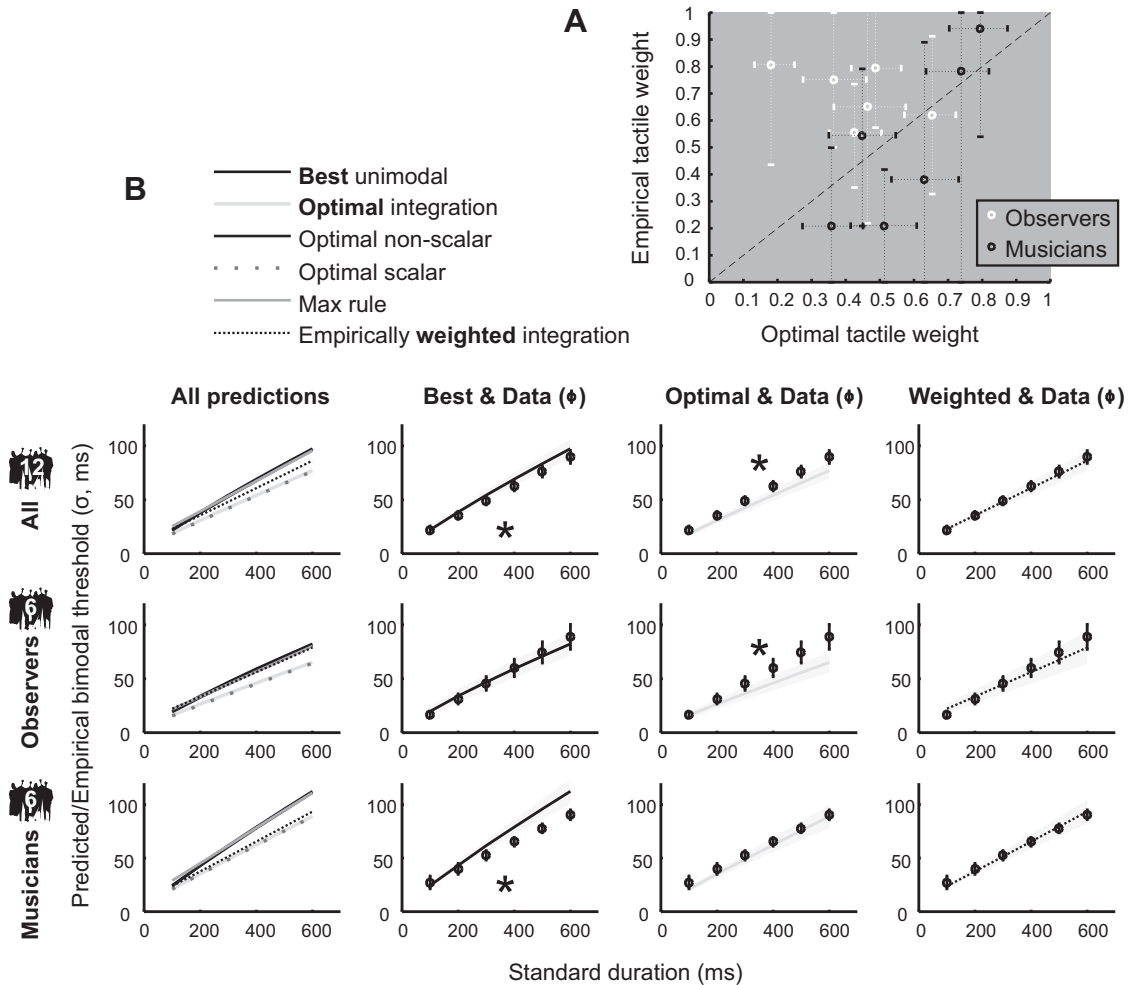


Figure 4. Group average results of Experiment 2. **A:** Scatter plot of empirical versus optimal (predicted) tactile weights for both musicians and psychophysical observers. Error bars show 95% bootstrap confidence intervals. Musicians appear to select near optimal weights. **B:** Comparison of predictions for various models against mean thresholds from the bimodal conditions. Data are shown for all participants (top) or separately for psychophysical observers (middle) and musicians (bottom). Model predictions are shown in the left hand set of graphs. Predictions from selected models are then repeated with their standard errors (gray regions) alongside bimodal data (error bars show standard errors) for (from left to right) the best unimodal prediction, the optimally weighted integration prediction, and the empirically weighted integration prediction (based on empirical weights from Panel A). Asterisks (*) show significant differences between models and data.

optimally weighted integration model predictions and bimodal data for psychophysical observers, $F(1, 5) = 27.95$, $p = .003$, particularly at longer durations (Model vs. Data \times Duration Interaction), $F(5, 25) = 21.66$, $p = .005$, but not for musicians ($F_s \leq 1.0$).

Finally, the rightmost row of graphs show model predictions alongside bimodal data for predictions based on the weighted combination of cues, but with weights estimated empirically and (particularly for the psychophysical observers) tending to deviate from the optimal choice. In this case, it is clear that model and data are in close agreement, both for the overall sample and when broken down by musical subgroup, with ANOVA showing no effects involving the model versus data contrast (all $p_s > 0.1$).

The comparison of each model's predictions to bimodal data for our entire sample (and to a lesser extent for each subgroup) provides a sense of how accurately the models perform on average. However, this may disguise problems predicting performance for each individual observer. Given the larger number of trials (and subsequent data cleaning) used in Experiment 2 compared to Experiment 1, individual estimates were more reliable, so we also examined predictions for each participant separately. These are presented for our three most distinct models in Figure 5, expressed as differences between model predictions and bimodal thresholds. Figure 5 also summarizes occasions where bootstrap contrasts were found to indicate significant differences between models and data for each par-

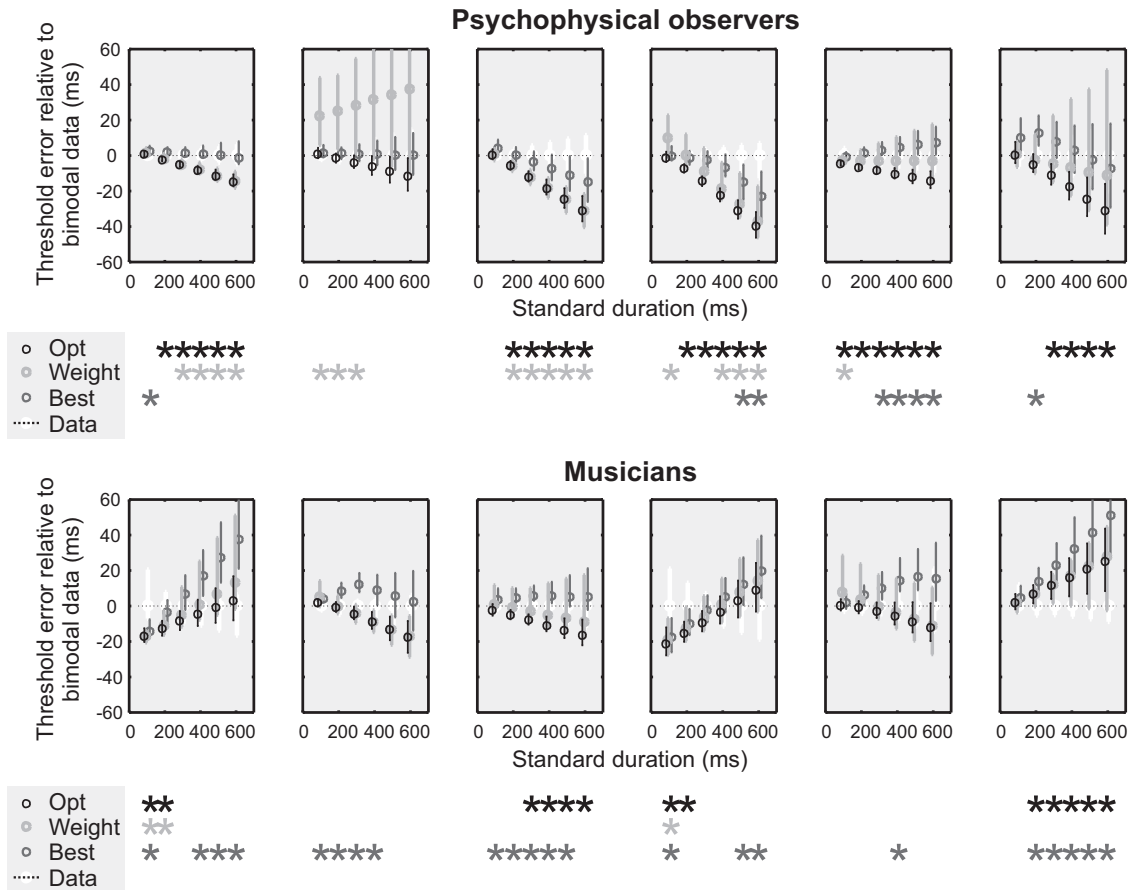


Figure 5. Individual data for Experiment 2, shown for all psychophysical observers (top) and musicians (bottom). Model predictions are shown as deviations from thresholds observed in the bimodal conditions. Opt = optimally weighted integration prediction; Weight = weighted average integration prediction; Best = best unimodal prediction. Error bars show 95% bootstrap confidence intervals. Asterisks (*) below plots denote significant differences between models and data.

participant (uncorrected for multiple comparisons). It is apparent that although the weighted integration model performed well for (sub)group-averaged data, it was not particularly accurate at the individual level, and although it was rejected less often for our sample than alternative models (in which participants are assumed to either rely on their best unimodal estimate or to combine cues optimally) this was partly because its predictions are estimated less precisely.

In summary, Experiment 2 generated average bimodal data that differed significantly from predictions assuming a simple reliance on the best sensory modality, and from the predictions of MLE cue combination, but which matched well with a weighted averaging process in which weights might be selected suboptimally. However, even empirically weighted averaging struggled to predict thresholds for each participant considered individually.

General Discussion

We ran two experiments in which participants (with or without high levels of musical training) judged the durations of subsecond

intervals that could be filled with either vibrotactile, visual, or vibrotactile and visual signals. Performance in unimodal conditions was used to predict bimodal performance via various models of sensory and decisional processes. In Experiment 1, participants performed better than predicted by assuming that they could use just one of the available signals, but worse than predicted assuming they had based decisions on an optimally weighted summation of initially independent sensory estimates. Hence neither model received strong support.

In Experiment 2 we were able to explore additional models and, in particular, a weighted-integration model which combined cues using empirical weights suggested by PSE differences when the standard stimulus contained a conflict. This model which, unlike MLE integration, did not imply that participants had accurate knowledge of their own unimodal precision, predicted data well, at least at the group-average level. Although we cannot test it formally against data from Experiment 1, it seems qualitatively consistent with those data too. However, it may simply be that some people are near optimal, whereas others fail to integrate information via any form of averaging. To the extent that weighted

averaging can summarize our data, the suboptimal selection of weights did not seem to reflect overall performance (our psychophysical observers in Experiment 2 often had lower thresholds than our musicians) suggesting that the musician's life experiences might have caused them to select more optimal weights. However, this assertion should be treated with caution given the small sample in Experiment 2 and the failure of the sample of musicians in Experiment 1 to achieve optimal integration.

Several previous studies have used methods similar to ours to assess the integration of bimodal duration cues, although they have generally examined only one or two models of bimodal performance, which may create a false sense of certainty regarding how well particular accounts have fared. Most such studies seem broadly consistent with our results. For example, both Shi et al. (2013) and Tomassini et al. (2011) report suboptimal integration (at least based on precision measures) for tasks involving a tactile component. However, two audio-visual experiments stand out for generating results that appear to favor either no integration at all (Eijkman & Vendrik, 1965) or optimal integration (Hartcher-O'Brien et al., 2014). It is not clear which procedural differences (e.g., duration, 1000 ms vs. 500 ms respectively; stimulus degradation, none vs. substantial, respectively) are critical in generating these contrasting results. However, we speculate that either the quality of the observers (visual weber fractions $\sim 10\%$ vs. $\sim 40\%$ respectively, both based on σ as a threshold measure) or the degree of spatial overlap between unimodal stimuli (none vs. complete, respectively) might be important. Our own experiments were somewhat intermediate in relation to both of these measures, with average visual weber fractions ranging from $\sim 45\%$ (nonmusicians, E1) to $\sim 17\%$ (psychophysical observers, E2) and a visual stimulus that was placed near, but not exactly over, the vibrotactile stimulus. Of course, both previous reports might be consistent with our own (and with each other) if selection of weights by participants was simply particularly unfortunate in Eijkman and Vendrik's (1965) work and particularly fortunate in Hartcher-O'Brien et al. (2014).

Our findings have clear implications for the current debate concerning the nature of the internal clock, and in particular for its unitary versus distributed character. Any form of improvement through averaging, even one based on suboptimal weights, implies the existence of two duration estimates subject to partially independent sources of noise. However, this leaves open several possibilities regarding where such noise might arise, and hence several plausible clock architectures. Two such possibilities are schematized in Figure 6. If we assume that the major source of noise in duration judgments comes from some kind of counting or temporal integration process (e.g., Taatgen, van Rijn, & Anderson, 2007), then our data would point toward the existence of two separate counters, one tactile and one visual. However, classical amodal clock models have tended to attribute most of the noise in duration perception tasks to memory processes, for example to noise arising during the conversion of an accumulated clock count into a stored memory (Gibbon, Church, & Meck, 1984).⁷ Hence another viable account of our data would be to posit just one counter, with transfer of the accumulated estimate into two (sensory specific) memory stores that each contribute independent noise. Of the two possibilities, the latter could be said to be less parsimonious, as it assumes dual independent memory stores in addition to early, initially independent, sensory processes, but we note that there is

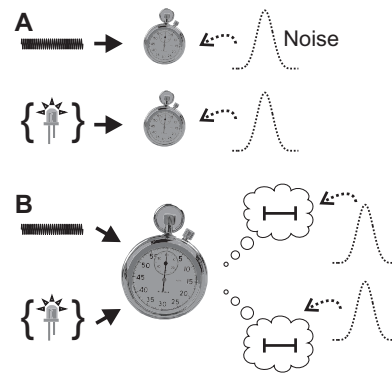


Figure 6. Schematic of two possible internal-clock architectures. A: Separate clocks exist for tactile and visual stimuli, each contributing independent noise, with estimates averaged prior to a comparison decision. B: A single clock receives both tactile and visual inputs, but contributes relatively little noise. Its output is stored in separate tactile and visual memory stores, with independent noise accrued at this stage (or later) before estimates are averaged for a comparison decision.

some evidence from selective interference experiments suggesting that separate short-term memory (STM) stores might be used when interval estimates arise from different sensory modalities (e.g., Rattat & Picard, 2012, but see Filippopoulos, Hallworth, Lee, & Wearden, 2013).

Our data are consistent with broader evidence suggesting separate clocks for separate modalities, such as the modality-specific duration adaptation effects alluded to in the introduction (Heron et al., 2012). However, we are at odds with other findings, such as performance correlations between auditory and visual timing tasks (Merchant et al., 2008) and transfer of training benefits from one modality to another (Bartolo & Merchant, 2009; Bratzke et al., 2012; Nagarajan et al., 1998). Of course, both of these findings might result from the sharing of some, but not all, clock components across modalities (or, in the former case, the existence of multiple clocks either built in similar ways, or depending upon some general property, such as neural efficiency).

In addition to empirical findings, there are some theoretical puzzles affecting a multiple clock account. Interval timing does not require a continuous signal, and it is not noticeably worse without one, as when we time empty intervals. What is going on in these situations? If it is the recruitment of an amodal clock, then why, with this clock available over and above sensory-specific ones, isn't filled interval timing consistently better than empty interval timing? In fact, some psychophysical evidence suggests it may be (e.g., Horr & Di Luca, 2015; Rammsayer, 2010).

Our apparent ease at converting interval estimates arising in different modalities also makes an amodal clock account appealing. However, a lifetime of experience might bring different clocks into alignment. More generally, our current findings favoring sense-specific clocks form part of a wider body of

⁷ This attribution seems to rest mainly on the assumption that the counter would be a Poisson noise source. Although not unreasonable, this is by no means certain.

evidence that speaks against single amodal internal stopwatch accounts, like scalar expectancy theory, which can be criticized on both theoretical grounds (e.g., this model achieves scalar timing via an ad hoc multiplicative constant, rather than by any feature of its basic architecture; see Staddon & Higa, 2006) and on empirical considerations (including that humans do not appear able to stop and restart their timer at will, as this model suggests, without suffering large drops in precision; Narkiewicz et al., 2015).

There are, of course, caveats to our preferred interpretation. Regarding the degree to which we obtained optimal integration, we have already noted that the spatial congruence of our stimuli was only approximate. This might be an important factor. More generally, we are not aware of any multimodal study that has gone very far toward promoting an ecologically valid scenario for the combination of interval timing stimuli. Furthermore, our findings only apply to the range of subsecond intervals that we made use of. Different mechanisms might apply for suprasedond timing (Lewis & Miall, 2003; Rammsayer, 1999).

Another concern regards the assumptions underlying our various model predictions. We noted in the introduction that it can be difficult to formulate accurate predictions when decisions reflect multiple sources of noise. This point was made as we are mindful of studies using reproduction tasks, where motor noise contributes toward interval estimates but might not itself be amenable to integration (Shi et al., 2013; Tomassini, Gori, Burr, Sandini, & Morrone, 2011). Our comparison task was free from motor noise. However, a similar issue might apply to decision noise, for example trial-by-trial variance in the placement of a decision criterion. Any such noise, affecting estimates in unisensory conditions must also apply in entirety to the bimodal condition, so predictions assuming purely sensory noise might overestimate the degree of bimodal improvement that is possible. We were able to generate predictions for models that partitioned sensory noise into scalar and nonscalar components in our second experiment, and optimal integration of just one or the other of these sources appeared to constitute an unlikely explanation of our data. However, we acknowledge that with sufficient inventiveness there might be other model-based accounts that could explain our data in addition to the one we have presented.

In a similar vein, the framework applied in this study attributes improvements in precision to mechanisms that are fully specified within simple models of psychophysical decisions. It is possible to conceive of higher-level factors that could affect precision, which these models do not consider. For example, perhaps participants try harder, or concentrate more, when there are two stimuli, feeling that they should really do better in this situation. Predicted performance improvements are certainly not so substantial that we can reject this kind of account. Nonetheless, specific models of cue combination are both elegant and quantitatively precise in a way that higher-level explanations are not, and to that extent seem to justify the attention they have received.

Finally, it is possible that the conclusions drawn from Experiment 2, where data were cleaned by fitting a generalized version of Weber's law, are overstated or biased, because scalar timing with an additional constant source of noise might be an oversimplification of how timing precision varies with stimulus duration. For example, some reports suggest a "dipper" function relating thresholds to duration for very brief intervals (Burr, Silva, Cicchini,

Banks, & Morrone, 2009). However, this is likely to be the result of a separate mechanism operating in the flutter/flicker fusion range. The near miss to Weber's law has received considerable empirical support (reviewed in Wearden & Lejeune, 2008) for typical subsecond-range interval timing, despite occasional deviant results (e.g., Kristofferson, 1980).

In conclusion, overall our experiments suggest that at least some people are able to extract and combine independent estimates of duration from simultaneously presented tactile and visual stimuli. However, average improvements in precision based on cue combination fall short of those predicted by an optimally weighted summation of initially independent sensory estimates. Hence our data suggest that participants do not generally have accurate knowledge of their own unimodal encoding precision. Regardless of whether all subjects perform integration, but suboptimally, or some integrate optimally and some not at all, our findings suggest that any realistic model of interval timing must incorporate sense-specific cognitive components that contribute a large part of the total variability in the precision of people's duration decisions. These components are subject to independent noise and are likely located at either the counter or memory stage of temporal information processing.

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